# Endogenous Transport Costs in International Trade

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# Abstract

In this paper we claim that distance alone is a poor proxy for international transport costs in empirical studies. We model a manufacturing and a transport sector and let the level of manufacturing exports determine the demand for transport services. Above a particular trade level, transport service suppliers nd it pro t-maximizing to invest in an advanced transport technology, which lowers their marginal costs and as a consequence, equilibrium transport prices. Transport costs thus vary with two characteristics: with the distance between two locations and with the endogenous decision to invest in a more e cient technology which is driven, in turn, by the bilateral export level. A simulation exercise reveals that ignoring the e ect of the investment decision on transport costs biases empirical results. The empirical estimations rely on newly collected transport price data from United Parcel Service (UPS). We apply an instrumental variable (IV) estimator to account for the endogeneity of the investment decision. Our results con rm that transport prices are in uenced by both the distance and the level of exports between two countries. We nd that trade partners with 10% more exports enjoy 0.8% lower transport prices.

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## 1. Introduction

Falling cross-border transaction costs have stimulated an unprecedented increase in cross-border transactions of all kinds. This rise in international activities has been impressive enough to label the past two decades nothing short of an era of globalization. The falling costs of cross-border transactions are thereby at the same time the source of the globalization process and its result. Increased cross-border transactions do not only stem from transport cost reductions, they also boost investments in the infrastructure of international trade, which induce further cost cuts. In the light of the importance of this phenomenon, the scarcity of economic studies that address the role of infrastructure investments in lowering transport costs and in stimulating trade, is surprising. In this paper, we put the infrastructure of international trade and the transport sector where the investment decisions are taken in the focus of the analysis.

We start with the observation that the costs of transporting goods between two countries seem to vary not only with respect to the distance between them. While most Asian economies, rst and foremost, China, trade high volumes at moderate transport prices with the United States and the European Union (EU), many African economies trade rather moderate volumes at high transport prices despite of their more favorable geographic location. Hence, there must be more than just distance a ecting transport costs. The recent literature has stressed that transport costs di er systematically with the market structure of the transport sector (Hummels et al., 2009), with bilateral trade imbalances (Behrens and Picard, 2011 and Jonkeren et al., 2011), and with port e ciencies (Clark et al., 2004 and Blonigen and Wilson, 2008). Complementing these ndings, we argue that bilateral trade levels are an important, yet largely neglected driver of the di erences.

To back the argument, we develop a theoretical framework that explicitly models a manufacturing and a transport sector and focuses on the investment decision of transport service suppliers. Transport service suppliers can choose between two route-speci c technologies: (i) a low xed costs / high variable costs technology and (ii) a high xed costs / low variable costs technology. This choice is motivated by the fact that increasing

in the trade imbalance. Firms in the net exporting country face higher, rms in the net importing country face lower transport costs if transport rms optimally set prices for the return journey. This price wedge works against the agglomeration forces. Thus, endogenous transport prices mitigate the separation of countries in an industrial core and an agricultural periphery which is so prominent in the models using iceberg transport costs.

Starting with Clark et al. (2004), a number of empirical studies have identied economies of scale as a determinant of transport costs (see e.g. Wilmsmeier et al., 2006, Martínez-Zarzoso and Wilmsmeier, 2010 and Pomfret and Sourdin, 2010). Clark et al. (2004) nd higher transport costs on routes with lower trade volumes. Assuming that any e ect of country size on transport costs goes through trade volumes, they use GDP as an instrumental variable (IV) for trade volumes. The negative e ect of trade volumes on transport prices becomes more pronounced when exports are instrumented, suggesting that failing to account for the endogeneity of exports understates their impact. Using the gap between c.i.f and f.o.b values of Australian imports as a measure of transport costs, Pomfret and Sourdin (2010) show that country size explains some of the variation in transport costs along with distance, the weight of the product, and the institutional quality of the exporting and/or the importing country. Once imports are used as a regressor instead of GDP to approximate country size, the signi cantly negative e ect on trade costs becomes larger and more robust. Clark et al. (2004) and Blonigen and Wilson (2008) make a reference to technology. They argue that di ering port e ciencies explain the country-speci c part of transport cost variations whereas variables such as distance, trade imbalances and product weight are introduced to capture the bilateral transport cost determinants.

transport sector. The decision to supply transport services thereby involves a decision about an investment in a particular transport technology.

Although we focus on the determinants of transport prices, the analysis naturally relates also to the literature that deals with the correct speci cation of the gravity equation. Endogeneity problems in gravity equations have provoked lengthy discussions in the trade literature of the past decade. Nearly all of the typically employed variables have been surmised to simultaneously in uence trade, and, be in uenced by trade. The usual suspects include national incomes (Frankel and Romer, 1999) and Free Trade Agreements (FTAs) (see e.g. Baier and Bergstrand, 2004 and Egger et al., 2010). Mostly approximated by time-invariant distance, transport costs have, by contrast, been perceived as exogenous and even served as an instrumental variable for trade assuming their orthogonality to other gravity variables (Frankel and Romer, 1999).

A notable exception is Rudolph (2009) who argues that scale economies leading to falling average costs arise in the presence of xed costs in the transport sector. Not accounting for the endogenous impact of trade on transport costs biases the coe cients of traditionally estimated gravity equations. Rudolph (2010) applies a simultaneous equation model to jointly estimate trade and transport costs, the latter being approximated by the trade volume within the respective economies relative to the trade volume between them. He presents two ndings: rst, trade levels and transport costs are simultaneously determined. Second, ignoring the simultaneity results in overestimating the impact of transport cost proxies on trade. In order to provide a more reliable estimate

sector. We account for the evidence they present by modeling an oligopolistic sector which invests in route-speci c instead of country-speci c infrastructure. For clarity, we abstract from Hummels et al. (2009)'s nding of di erent degrees of competition and focus instead on the technology choice.

#### 3. Theoretical Framework

In this section, we develop a two-sector model that formalizes the argument that bilateral trade levels and bilateral transport prices are jointly determined. The model consists of an oligopolistic transport sector, T, (with a xed number of rms) which produces a homogenous transport service for each route and of a monopolistically competitive manufacturing sector, M, where exporting rms face per-unit transport costs. Prices are determined in equilibrium where the units of o ered transport services equal the units of goods from the manufacturing sector that need transportation to a foreign country. To t the structure of the transport sector, we model the manufacturing sector based on a Melitz and Ottaviano (2008)-framework with a quasi-linear demand structure and additive transport costs. We choose the simplest possible set-up with labor as the only factor of production. There are  $L_j$  individuals in economy j, each o ering one unit of (homogenous) labor.

## 3.1. The Manufacturing Sector

The manufacturing sector, M, comprises N heterogenous rms that engage in monopolistic competition. Firms set prices depending on their marginal costs and decide about their export participation. Marginal costs depend on the rm-speci c productivity level that is drawn independently at market entry from a common distribution. This rm-speci c productivity is the only primary source of rm heterogeneity. The other, secondary source of heterogeneity, is the rm's export status which directly results from the heterogeneity with respect to productivity. In this static framework, consumers spend their complete income on the consumption of the goods produced in the manufacturing sector.

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## Consumers

Following Melitz and Ottaviano (2008), preferences of a representative individual from country j are described by a quadratic utility function,

$$U_{j} = q_{ij}^{c}(0) + \begin{bmatrix} Z \\ m^{2} \\ m^{2} \end{bmatrix} q_{ij}^{c}(m)dm = \frac{1}{2} \begin{bmatrix} Z \\ m^{2} \\ m^{2} \end{bmatrix} (q_{ij}^{c}(m))^{2}dm = \frac{1}{2} \begin{bmatrix} Z \\ m^{2} \\ m^{2} \end{bmatrix} (q_{ij}^{c}(m)dm = \frac{1}{2} \end{bmatrix} (1)$$

where  $q_{ij}^c(0)$  and  $q_{ij}^c(m)$  refer to the individual consumption of the numeraire and the di erentiated good, m. The rst index, i, refers to the country where the production of the di erentiated good, m, takes place. The second index, refers to the home country of the consumer. and indicate the degree of substitutability between the di erentiated varieties and the numeraire, governs the degree of di erentiation between the varieties. The inverse demand function is given by

$$p_{ij}(m) = q_{ij}^{c}(m) Q_{ij}^{c};$$
 (2)

where  $Q_{ij}^c = \frac{R}{m_{2}} q_{ij}^c(m)$ dm. With  $q_{ij} = L_j q_{ij}^c$  and  $q_{ij}^c > 0$ , we obtain the subset of produced varieties which satis es

$$p_{ij}(m) = \frac{1}{N_j + ( + N_j p_j)};$$
 (3)

where N denotes the number of rms and  $p_j$  the average price in country j with  $p_j = 1=N_m^P p_{ij}(m)$ . The consumer price,  $p_{ij}(m)$ , includes the per-unit transport costs,  $p_{ij}(m) = p_i(m) + t_{ij}$ , if the good is imported (j  $\in$  i).

# Producers

We assume that product di erentiation is costless which guarantees that each good m is produced by only one rm. Firms maximize pro ts,

$$_{ij}(m) = q_{ij}(m)(p_{ij}(m) c_i(m) t_{ij})$$
 (4)

for the foreign market (i  $\in$  j) and for the domestic market (i = j and t<sub>ii</sub> = 0) separately. While products enter symmetrically in the consumption bundle, we keep the rm index m because rms di er with respect to their productivity level. Firm-speci c productivity levels translate into rm-speci c marginal costs c<sub>i</sub>(m), rm-speci c prices p<sub>ij</sub>(m) and rm-speci c output levels q<sub>j</sub>(m). Using the residual demand from (2), rms obtain their output function as

$$q_{ij}(m) = \frac{L_j}{m} (p_{ij}(m) - c_i(m) - t_{ij}):$$
 (5)

A rm stays in the domestic and enters a foreign market if its price at least equals its marginal costs,  $p_{ij}(m) = c_i(m) + t_{ij}$ . We denote the maximum marginal costs for rms from country i to be active in market j as

of the maximum costs in country j,  $c_j$ , and the number of rms from country i, N<sub>i</sub>. The number of rms from country i that are active in country j can be expressed as the product of the share of exporters in the number of rms in i, N<sub>ij</sub> =  $\frac{G(c_{ij})}{G(c_{i})}$  N<sub>i</sub> =  $\frac{c_j}{c_i}$  N<sub>i</sub>.

In Appendix A.1, we show that the transport costs a ect the trade level negatively, i.e. that the partial derivative  $\frac{@Q}{@\sharp} < 0$ . Considering that exports are declared net of transport costs, we next obtain the total bilateral export value by aggregating each rm's export sales,  $r_{ij}^{fob}(m) = p_{ij}^{fob}(m)q_{j}(m)$ , over all exporters from i to j,

$$EX_{ij} = N_{i} \frac{L_{j}}{4} \int_{0}^{Z_{i}} c_{j}^{2} c_{i}^{2}(m) + 4t_{ij}^{2} - 4c_{j}t_{ij} g(c_{i}(m))dc_{i}(m)$$

$$= \frac{1}{4} N_{i}L_{j}c_{i} - \frac{2}{2+}(c_{j} - t_{ij})^{2+} + 3t_{ij}^{2} - 2c_{j}t_{ij} (c_{j} - t_{ij}) :$$

$$= \frac{1}{4} \sum_{\substack{i = 1 \\ i =$$

Equation (8) shows that the aggregate bilateral export values are characterized by a gravity-type relation where the two country sizes, N<sub>i</sub> and L<sub>j</sub>, a ect exports positively, while the transport costs,  $t_{ij}$ , a ect them negatively (since  $c_j = 2t_{ij}$ ). Furthermore, exports rise in the minimum (and therefore average) productivity of the home country f (1= $c_j$ ) and fall in the productivity of the partner country f (1= $c_j$ ). As a result of the additive transport costs, the partner country's productivity, (1= $c_j$ ), is strongly interlinked with the transport costs between the two countries,  $t_{ij}$ .

# 3.2. The Transport Sector

As the transport sector typically consists of a few, large companies, we impose an oligopolistic market structure. We assume that transport is a homogenous service. Consequently, exporting rms will base their decision for a particular transport service supplier entirely on cost considerations. To keep the model simple and to focus on di erences in the aggregate pattern of transport costs between two countries, we model the transport sector as consisting ofn<sup>T</sup> symmetric rms.<sup>2</sup> In a world with I exporting and J importing countries, I J transport routes exist. We assume that each transport rm

<sup>&</sup>lt;sup>2</sup>Imposing symmetry does not a ect our main argument while it simpli es the analysis considerably.

serves each route. The total number of transport  $rms, n^{T}$ , is exogenously given<sup>3</sup>.

Transport rms choose their transport technology when starting to service a particular route. Like Yeaple (2005) and Bustos (2011), we simplify this choice by assuming that there are just two possible cost structures to choose from: technology with low variable costs,  $a^L$ , and high xed costs,  $f^L$ , and technology H with high variable costs,  $a^H$ , and low xed costs,  $f^H$ , i.e.  $a^L < a^H$  and  $f^L > f^H$ 

with " =  $\frac{@Q_{ij}}{@t_{ij}} \frac{t_{ij}}{Q_{ij}}$  as the price elasticity of demand. Output, i.e. the supply of transport services, increases in the transport price  $t_{ij}$ , and the export quantity,  $Q_{ij}$ , of the manufacturing sector. With demand (as given in (7)) strictly falling and supply (as given in (11)) strictly rising in the transport price,  $t_{ij}$ , there exists exactly one transport price level that clears the market for transport services. Equation (11) also shows that the output of a transport service supplier is negatively a ected by the variable costs of supplying the service,  $a_{ij}$ .

The second equation in (11) uses the fact that transport rms are symmetric by assumption and that the transport service market must be cleared in equilibrium, hence  $Q_{ij}(t) = {P \atop 1}^{n^{T}} q_{ij}(t_{ij}) = n^{T} q_{ij}$ . Solving the supply equation (11) for the transport price,  $t_{ij}$ , yields the price as a function of the rms' costs,  $a_{ij}$ , the number of rms,  $n^{T}$ , and the demand elasticity, ",

$$t_{ij} = \frac{"n^{T}}{"n^{T} \ 1} a_{ij}$$
 (12)

Knowing that in a symmetric equilibrium every rm serves  $Q_{ij} = n^T$  of the demand, we can rewrite the pro ts from (10) as

$$_{ij} = (t_{ij} \quad a_{ij})q_{j} \quad f = \left(\underbrace{t_{ij}}_{_{ij}} \underbrace{a_{ij}}_{_{ij}}\right) \frac{Q_{ij}}{n^{T}} \quad f;$$
(13)

where we de ne the mark-up  $_{ij}$  as  $t_{ij}$   $a_{ij}$ . With this outline, we can now study the incentive to invest in a variable costs saving transport technology for the route between country i and j. Equation (14) uses (12) to show that the variable pro ts,  $_{ij}^{var}$ , generated on route ij increase as the marginal costs of shipping between these two countries fall,

$$\frac{d_{ij}^{var}}{da_{ij}} = \frac{@_{ij}}{@_{aj}} \frac{Q_{ij}}{n^{T}} + \frac{@_{aj}}{@_{aj}} \frac{ij}{n^{T}}$$

$$= B \frac{Q_{ij}}{n^{T}} \frac{1}{"n^{T} - 1} < 0;$$
(14)

where B 1  $\frac{(1+)[2\mathfrak{E}_{j} \quad (2+)t_{ij}]}{(\mathfrak{E}_{j} \quad t_{ij})[\mathfrak{E}_{j} \quad (2+)t_{ij}]} \frac{"n^{\mathsf{T}}}{"n^{\mathsf{T}} \quad 1} = 1 \frac{"}{a_{ij}} < 0$  if the price elasticity of demand " is not to low, i.e. if " > a <sub>ij</sub>. This holds if the transport costs are not too low, since

the price elasticity of demand rises with transport costs<sup>6</sup> In the following, we assume that the price elasticity of demand for shipping is su ciently high to ensure the negative relationship. Note, that there is a trade-o between lower mark-ups and larger demand following the cost reduction. Since the second e ect outweighs the rst, pro ts increase with falling costs. For our argument most important, equation (14) states that the pro t-rising e ect of investing in advanced technologies increases in the export volume,  $Q_{ij}$ , from the manufacturing sector of country i. Thus, routes on which large volumes of goods are traded generate more additional pro ts if the variable costs of transportation,  $a_{ij}$ , fall.

The comparison of pro ts guides the rm's decision of investing in one of the two available technologies. Transport suppliers decide to invest in the advanced technology if the lower marginal costs generate su ciently high variable pro ts to make up for the higher xed costs. The discussion above reveals that this is more likely for transport routes with high trade volumes,  $Q_{ij}$ ,

$$d_{ij} > 0 ! \frac{1}{n^{T}} (t_{ij}^{L} a_{ij}^{L}) Q_{ij}^{L} (t_{ij}^{H} a_{ij}^{H}) Q_{ij}^{H} > f^{L} f^{H}:$$
(15)

Routes that generate more additional variable pro ts are more likely to jump the additional xed costs hurdle  $f^L f^H$ . As argued above, the large trade volume routes create the largest additional variable pro ts. Hence, on these routes the introduction of the low variable costs technology is more likely. Since the technology choice depends on the trade volume, we expect lower transport prices on routes with large trade volumes. In turn, the technology choice a ects the marginal costs and therefore the transport prices,

$$t_{ij}^{I} = \stackrel{8}{\underset{ij}{\stackrel{}{\sim}}} \frac{n^{T} a_{ij}^{L}}{n^{T} 1 = "L} \quad \text{for a high trade volume}$$

$$\stackrel{1}{\underset{ij}{\stackrel{}{\sim}}} \frac{n^{T} a_{ij}^{H}}{n^{T} 1 = "H} \quad \text{for a low trade volume:}$$
(16)

Equation (16) shows that the transport costs for rms from the manufacturing sector

<sup>&</sup>lt;sup>6</sup>See Appendix A.2 for the derivation of this result.

di er for routes of similar distance and similar other characteristics if the chosen technology di ers. This implies that transport prices which are set in the transport sector eventually depend on export volumes decided on in the manufacturing sector, which, in turn, depend on transport prices. The mutual dependance of exports as given by equation (8) and transport prices as given by equation (16) reveals that both variables are, in fact, jointly determined.

4. Estimating Transport Costs: An Illustrative Example

porting and 61 importing countries<sup>7</sup>, from which we calculate a distance matrix between any two of these countries.<sup>8</sup> We draw an arbitrary size (GDP<sub>i</sub> > 0; GDP<sub>j</sub> > 0) for each of these countries from a uniform distribution with mean 500 and a country-speci c marginal cost threshold which corresponds to**¢**  country i to country j

above imply that investments into the advanced technology take place on about 15% of all trade routes. In the continuous version of the investment decision, we assume  $I_{ij} = 2(EX_{ij} = 10)^{0:15}$ . The parameters are chosen to match the maxima and minima of

	Omitting	Proxy	Proxy	Proxy	IV	IV
	I ij	top 150	top 250	top 350	I ij	EX ij
Depender	nt variable: Bila	ateral transport of	costs t <sub>ij</sub>			
dist <sub>ij</sub>	0.275 (0.0148)	0.209 (0.0101)	0.206 (0.00813)	0.214 (0.00898)	0.200 (0.00453)	0.229 (0.0116)
gdp	1.000 (0.0105)	1.000 (0.00707)	1.000 (0.00571)	1.000 (0.00633)	1.000 (0.00312)	1.000 (0.00817)
top 150		-1.054 (0.0232)				
top 250			-0.961 (0.0150)			
top 350				-0.788 (0.0145)		
l ij					-1.000 (0.0176)	
EX <sub>ij</sub>						-0.165 (0.00474)
$R^2$	0.84	0.92	0.95	0.94	0.99	0.90

Table 1: Addressing the Omitted Variable Bias in the Transport Cost Estimation: the Discrete Case

Table 2: Addressing the Omitted Variable Bias in the Transport Cost Estimation: the Continuous Case

	Omitting I <sub>ij</sub>	Proxy top 150	Proxy top 250	Proxy top 350	IV I <sub>ij</sub>	IV EX <sub>ij</sub>
Depender	nt variable: Bila	teral transport o	osts t <sub>ij</sub>			
dist <sub>ij</sub>	0.230 (0.00906)	0.209 (0.00852)	0.208 (0.00821)	0.207 (0.00793)	0.200 (0.00435)	0.198 (0.00434)
gdp	1.000 (0.00640)	1.000 (0.00597)	1.000 (0.00577)	1.000 (0.00558)	1.000 (0.00309)	1.000 (0.00308)
top 150		-0.328 (0.0196)		. ,	. ,	
top 250			-0.315 (0.0151)			
top 350			. ,	-0.309 (0.0128)		
l <sub>ij</sub>				· · /	-1.000 (0.0141)	
EX ij					. ,	-0.138 (0.00182)
R <sup>2</sup>	0.93	0.94	0.94	0.95	0.98	0.98

Note: Standard errors in parentheses.

Source: Own calculations.

350 export routes come close to the coe cients of the true model. In the continuous case, using the instrumented export level, however, works better. The overidenti cation test

rejects the validity of the instruments in less than 10% of the repetitions with the continuous investment indicator, while it rejects the validity of the instruments for nearly 50% of the repetitions with the discrete investment indicator. Because the instruments are exogenous by construction, this result demonstrates how sensitive the overidenti cation test reacts to the nature of the omitted investment variable, i.e. whether it is aisare working with real data and (ii) because poisson regressions allow for a correct treatment of zero trade ows. Even though the number of zeros is very low in this OECD countriescentered sample, the poisson estimations provide an important robustness check.

# 5.1. Data

Bilateral transport costs are di cult to measure. <sup>11</sup> We have built a new data set by collecting information from UPS on the costs of shipping a10kg package per express delivery between two countries. Transport prices for 2010 are available for 61 countries. In cases where di erent prices apply to di erent regions of one country, we take the prices of the region to which the most populated city belongs.

We analyze the transport prices charged on di erent routes together with bilateral trade data. The OECD ICTS database provides bilateral trade data in US\$ for 30 OECD countries (all member states as of 2009) with partner countries worldwide. The latest available year for which the data is complete, is currently 2009. We select the 61 trade partners for which we were also able to gather information on transport prices. If we had full information, we would have a data set containing 30

## Figure 1: UPS Transport Prices and Exports

# 5.2. Main Results

We start with estimating transport prices by using a single equation approach. In order to make our results match the predominant number of empirical studies, we primarily report OLS estimates. Table 3 shows the estimation of transport prices as a function of distance and GDP per capita in Column 1. The results indicate that rms set higher prices on more distant routes. The impact of distance on transport prices is, however, moderate. Transporting goods between countries which are 10% farther away from each other is 1.43% more expensive, on average. This result is in line with Clark et al. (2004) and other empirical estimations of transport cost equations that do not control for route-speci c investments. Other marginal costs, captured by per capita GDP, have a similar e ect.

	Omitting I <sub>ij</sub>	Proxy top 150	Proxy top 250	Proxy top 350	IV EX ij
Dependent varia	able: t <sub>ij</sub>				
dist <sub>ij</sub>	0.143*** (0.018)	0.121*** (0.017)	0.117*** (0.016)	0.112*** (0.017)	0.083*** (0.017)
gdp	0.132*** (0.009)	0.145*** (0.009)	0.150*** (0.009)	0.148*** (0.009)	0.163*** (0.009)
top 150		-0.383*** (0.049)			
top 250			-0.345*** (0.037)		
top 350				-0.293*** (0.030)	
EX ij					-0.076*** (0.008)
N R <sup>2</sup>	1,740 0.151	1,740 0.205	1,740 0.220	1,740 0.214	1,739 0.232
Endog. test p-val.					12.585 0.000
Hansen J p-val.					1.241 0.265
Underid. test p-val.					57.35 0.000
P val. Weak id. test p-val.					1730.70 0.000

Table 3:	OLS Estir	nation of	Transport	Prices
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Note: Cluster-robust standard errors in parentheses with signi cance at the \*\*\* p<0.01, \*\* p<0.05, \* p<0.1 level. Source: Own calculations.

Since we cannot observe investments in the transport sector, we follow the strategy outlined in Section 4 and let the R<sup>2</sup>s guide our decision about the most appropriate speci cation. In Columnslet to0

	Omitting	Proxy	Proxy	Proxy	IV
	Lij	top 150	top 250	top 350	EX ij
Depender	nt variable: t <sub>ij</sub>				
dist <sub>ii</sub>	0.131***	0.113***	0.109***	0.104***	0.088***
	(0.018)	(0.017)	(0.016)	(0.017)	(0.018)
gdp	0.132***	0.142***	0.146***	0.145***	0.152***
	(0.007)	(0.007)	(0.007)	(0.007)	(0.006)
top 150	. ,	-0.391***	. ,	. ,	, , , , , , , , , , , , , , , , , , ,
•		(0.046)			
top 250		. ,	-0.347***		
•			(0.037)		
top 350				-0.289***	
•				(0.030)	
EX ij				· · /	-0.054***
					(0.007)
N	1,740	1,740	1,740	1,740	1,739
R <sup>2</sup>	0.126	0.171	0.185	0.181	0.198

Table 4: Poisson Estimation of Transport Prices

Note: the single equation estimations are reported with cluster-robust standard errors in parentheses with significance at the \*\*\* p<0.01, \*\* p<0.05, \* p<0.1 level.

Source: Own calculations.

strong predictor of transport prices, even though there is no exclusive relationship as suggested by the gravity literature which often relies entirely on distance to approximate transport costs. Instead, the infrastructure investments, which we approximate by dummy variables for the most frequented export routes and by the bilateral export level are important transport price determinants as well. Including the investment proxies increases the R<sup>2</sup> also in the poisson regressions. Again, it is highest in Column 5, where we include and instrument the bilateral export level. Moreover, the distance coe cient falls which we, based on the theoretical considerations and the simulation, interpret as bias correction.

# 5.3. Robustness Checks

Transport service suppliers like the UPS do not o er transport for all kinds of goods. Certain raw materials, for example, need a very speci c infrastructure, like pipelines, on which cross-border transport crucially hinges. To account for the fact that the collected UPS transport prices may not apply to all goods, we repeat the estimations in Table

## 6. Transport Sector Investments and Globalization

The empirical results reveal that distance a ects transport prices positively. Yet, geography is not a destiny although it poses a challenge to policy-makers. Our results point to a strong price reducing e ect of bilateral trade values, which we interpret as stemming from investments in the transport sector. These investments in new, often large-scale trade-enhancing transport technologies result in important cost savings. Falling transport costs, i.e. falling costs of trade are the drivers of globalization. Even though the debate on the distance puzzle or the missing globalization puzzle (Coe et al., 2007) mentions that new technologies in transport bring down the costs of trade, in empirical applications such technology changes do not play a role. Instead, transport costs are modeled as iceberg costs which increase in distance. Consequently, globalization is searched for in the distance coe cient.

We argue that nding the source of falling trade costs requires augmenting international trade models by a transport sector, where the transport prices are actually set. When setting up the transport sector, we explicitly model rms that face a technology choice. Firms choose route-speci c technologies to maximize pro ts on each route. Modeling both the manufacturing goods sector and the transport sector enables us to show that specifying transport costs as a mere function of distance and distance-related variables misses an important point and creates an omitted variable problem in empirical applications. Since actual investments are often unobservable, the regressions require a proxy variable as a regressor.

From a policy perspective, there are at least three arguments that support the inclusion of transport sector investments in the analysis of international trade: it helps understanding (i) the driving force of globalization, (ii) the distance puzzle, and (iii) trade and development.

First, our model has the nice reinforcing feature of increasing trade leading to falling

dummies. We obtain qualitatively the same results with a slightly higher distance and per capita GDP coe cient. These additional results are not reported here but will be made available upon request.

transport prices which again stimulate trade. We think that the globalization process is well explained by such an investment-induced fall in transport prices. Seen like this the driving force of globalization is endogenous and the outcome of a pro t-maximizing behavior in the transport sector. The empirical results support this view. Transport prices are driven down by trade levels. Unfortunately, we do not have time series data and take all variation from the cross-section but if the production function of the transport sector did not miss anything important, infrastructure investments are a main source of di erences in transport prices on routes with similar distances. And, they are easily included as a regressor in the transport price equation.

Second, the study relates to the distance puzzle in two rather di erent ways. On the one hand, we observe that the omitted variable biases the distance coe cient in the transport price equation upwards which might contribute to the biased distance coe cient in gravity equations. On the other hand, we argue that if estimated correctly, the distance coe cient should be una ected anyway by technological changes that do not in uence the distance elasticity of transport costs. If this elasticity is systematically reduced, we are probably confronted with an investment function which is not only a ected by the trade level but also by distance. That might be the case if the technology choice includes the decision about di erent modes (air, sea, land, rail, pipeline) of transport which we have abstracted from in this study.

Third, the circular causality of transport prices and trade levels is very important for developing countries. Any change that leads to increasing trade has the potential to reduce transport prices and any transport price reduction increases trade levels. A regional trade agreement for instance might induce more trade and therefore lower transport prices which reinforce trade integration even more. Better investment conditions in the transport sector might induce investments in modern technologies and lower transport prices, thereby increasing the level of trade. One can think of many other policy measures that could help relatively closed economies to start a virtuous circle of lower transport costs and larger trade.

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## 7. Conclusions

Unlike most of the literature that assumes exogenously set, iceberg-type transport costs, this paper proposes marginal costs and prices in the transport sector to be endogenous and a ected by the bilateral export levels between two countries. By setting up a theoretical framework which comprises a manufacturing and a transport sector, we show that optimizing transport service suppliers invest in modern transport technology on highly frequented trade routes. The technology choice a ects transport prices via the marginal costs of supplying transport services between two locations. If this description of investment in the transport sector is correct, it is not su cient to approximate transport costs by distance and distance-related variables as done in the vast majority of empirical trade applications.

Using a constructed data set, we illustrate that the bias stemming from the omission of the investment decision in the transport sector, and its endogeneity to bilateral trade levels can be cured by using proxy variables and IV techniques. Employing a new data set which contains information on UPS transport prices, we detect an e ect from exports on transport prices. We nd that two countries with exports 10% above the average for all trade pairs enjoy 0.8% lower transport prices. This result adds to the discussion of the drivers of globalization, the distance puzzle and the development of economies which are currently lagging behind in terms of trade openness.

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# Appendix B. Country List

Australia	Finland	Ireland	New Zealand	Spain
Austria	France	Italy	Norway	Sweden
Belgium	Germany	Japan	Poland	Switzerland
Canada	Greece	Luxembourg	Portugal	Turkey
Czech Republic	Hungary	Mexico	Slovak Republic	United Kingdom
Denmark	Iceland	Netherlands	South Korea	United States

# Table B.1: List of Exporting Countries

# Table B.2: List of Importing Countries

Algeria	Croatia	Iceland	Netherlands	Slovenia
Argentina	Czech Republic	India	New Zealand	South Africa
Australia	Côte d'Ivoire	Indonesia	Nigeria	South Korea
Austria	Denmark	Ireland	Norway	Spain
Belgium	Egypt	Israel	Panama	Śweden
Brazil	Estonia	Italy	Peru	Switzerland
Bulgaria	Finland	Japan	Philippines	Thailand
Canada				

Variable	Mean	Std. Dev.	Min	Max
t <sub>ii</sub>	426.4566	180.4115	65.35735	989.058
dist <sub>ij</sub>	6444.817	5042.864	59.61723	19629.5
gdp	38277.33	18457.15	8720	84640
top 150	.0833333	.2764622	0	1
top 250	.1388889	.3459266	0	1
top 350	.1944444	.3958824	0	1
EX <sub>ij</sub> (in Tsd. US\$)	3760000	13200000	0.439	237000000
GDP <sub>i</sub> (in Mrd. US\$)	1360	2620	12.1	14100
GDP <sub>j</sub> (in Mrd. US\$)	879	1990	12.1	14100
Source: Own calculations.				

Table C.2: Descriptive Statistics: Real Data Set

Appendix D. Additional Empirical Results

		larde	l arde Sample		-	Small	Small Samole	
	(				(			
	Proxy	Proxy	Proxy	2	Proxy	Proxy	Proxy	2
	top 150	top 250	top 350	EX <sub>ij</sub>	top 150	top 250	top 350	EX <sub>ij</sub>
Dependent variable: t <sub>ij</sub>	ble: t <sub>ij</sub>							
dist <sub>ij</sub>	0.121***	0.117***	0.112***	0.084***	0.118***	0.116***	0.109***	0.079***
	(0.016)	(0.016)	(0.017)	(0.018)	(0.016)	(0.016)	(0.017)	(0.018)
gdp	0.144***	0.147***	0.145***	0.161***	0.145***	0.145***	0.144***	0.160***
	(0000)	(0.00)	(0.009)	(00.0)	(00.0)	(0.009)	(00.0)	(0.009)
top 150	-0.397*** (0.046)				-0.428*** (0.047)			
top 250		-0.340***				-0.348***		
		(0.037)				(0.035)		
top 350			-0.292***				-0.312***	
			(0.032)				(0.032)	
EX <sub>ij</sub>				-0.076***				-0.075***
				(0.008)				(0.008)
z	1,740	1,740	1,740	1,739	1,740	1,740	1,740	1,739
$R^2$	0.210	0.218	0.213	0.231	0.219	0.221	0.222	0.245
Endog. test				13.635				6.330
p-val.				0.000				0.012
Hansen J				1.273				1.618
p-val.				0.256				0.203
Underid. test				57.437				56.979
p-val.				0.000				0.000
Weak id. test				1753.73				1420.70
p-val.				0.000				0.000
Note: Cluster-robust standard errors in parentheses with signi cance at the *** p<0.01, ** p<0.05, * p<0.1 level	standard errors i	in parentheses wit	th signi cance at th	ie *** p<0.01, ** p<	<0.05, * p<0.1 leve	i.		

Table D.1: Robustness: Restricted Samples

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Source: Own calculations.

	Omitting I <sub>ij</sub>	Proxy top 20%	Proxy top 25%	Proxy top 30%	IV EX <sub>ij</sub>
Dependent varia	able: martc <sub>ijk</sub>				
dist <sub>ij</sub>	0.361*** (0.033)	0.309*** (0.033)	0.287*** (0.033)	0.308*** (0.031)	0.252*** (0.034)
gdp	-0.046* (0.026)	0.024 (0.027)	0.035 (0.027)	0.021 (0.027)	0.037 (0.029)
top 20%		-0.344*** (0.067)			
top 25%			-0.385*** (0.065)		
top 30%				-0.313*** (0.059)	
EX <sub>ij</sub>					-0.062*** (0.013)
Ν	6,754	6,754	6,754	6,754	6,754
R <sup>2</sup>	0.039	0.055	0.062	0.056	0.069
Endog. test					3.900
p-val.					0.048
Hansen J					1.605
p-val.					0.205
Underid. test					89.77
p-val.					0.000
Weak id. test					2598.91
p-val.					0.000

Table D.2: Robustness: Maritime Transport Costs

Note: Cluster-robust standard errors in parentheses with signi cance at the \*\*\* p<0.01, \*\* p<0.05, \* p<0.1 level. Source: Own calculations.

	Omitting I <sub>ij</sub>	Proxy top 150	Proxy top 250	Proxy top 350	IV EX <sub>ij</sub>
Dependent variab	le: t <sub>ij</sub>				
dist <sub>ij</sub>	0.115***	0.108***	0.107***	0.105***	0.086***
	(0.016)	(0.016)	(0.016)	(0.016)	(0.019)
gdp	0.138***	0.144***	0.147***	0.145***	0.162***
	(0.01)	(0.009)	(0.009)	(0.009)	(0.01)
top 150		-0.213***			
		(0.048)			
top 250			-0.198***		
			(0.037)		
top 350				-0.152***	
				(0.032)	
EX <sub>ij</sub>					-0.068***
					(0.015)
trade imbalance	-0.063***	-0.051***	-0.047***	-0.049***	-0.01
	(0.006)	(0.006)	(0.007)	(0.007)	(0.012)
N	1,737	1,737	1,737	1,737	1,737
R <sup>2</sup>	0.241	0.255	0.258	0.254	0.238
Endog. test					8.703
p-val.					0.003
Hansen J					1.181
p-val.					0.277
Underid. test					42.726
p-val.					0.000
Weak id. test					125.89
p-val.					0.000

Table D.3: Robustness: Trade Imbalances

igi р p level.

Source: Own calculations.